
Itô's lemma and applications

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1 Itô's early works on stochastic differential equations

- Differential equations determining a Markoff processes, Journal of Pan-Japan mathematical colloquium No.1077(1942) (original Japanese: Zenkoku Sizyo Sugaku Danwakai-si)
- Stochastic differential equations in a differentiable manifold, Nagoya Mathematical Journal vol.1, 35-47(1950)
- On stochastic differential equations, Memoirs of American Mathematical Society vol. 4,1-51 (1951)
- On a formula concerning stochastic differentials, Nagoya math. Journal vol.3, 55-65 (1951)

This is the cover sheet of
Pan - Japan mathematical colloquium
244 (1942)

全國紙上數學談話會

第244號

昭和十七年十一月二十日

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|----------------------------------|-------------|
| 1077. <i>Markoff</i> 過程ヲ定メル微分方程式 | 伊藤清(1352) |
| 1078. 一般タウバー型定理ノ理論的証明 | 深宮政範(1402) |
| 1079. <i>Wimann</i> ノ定理ニツイテ | 有馬喜八郎(1406) |

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This is the first page of
Ito's paper entitled "differential
equations which determine
Markov processes"

1077. Markoff 過程ヲ定メル微分方程式

伊藤 清(内閣統計局)

ハシガキ

(I) 有限個ノ可能ナ場合 a_1, a_2, \dots, a_m ヲ有シ, 自然
数ヲ径数トスル *simple Markoff process* x_1, x_2, \dots
ニ関シテ, 多クノ遷移確率ヲ考ヘルコトが出来ル. 例ヘバ
 $x_k = a_i$ ナル條件ノ下ニ於ケル $x_{k+1} = a_j$ ノ確率, 或ハ $x_k = a_i,$
 $x_{k+1} = a_j, \dots, x_{k+n} = a_{i+n}$ ナル條件ノ下ニ於ケル
 $x_{k+n+1} = a_{i+n+1}$ トナル確率等々. シカシ乍ラソレ等ハ結局
 $x_k = a_i$ ノ時ノ $x_{k+1} = a_j$ ナル確率 $p_{ij}^{(k)}$ ($k=1, 2, \dots, i, j$
 $= 1, 2, \dots, m$) ニ帰着セラレル. コレハ Kolmogoroff
ノ本^{(*)1}ニモ書イテアル. 以後コレヲ基本^{(*)1}ノ遷移確率ト呼
バク.

更ニ可能ナ場合ガ有限デナクトモ, 例ヘバ実数ヲ以テ
標識ヅケラレル時ニハ, 同ジコトガイヘルノハ云フマデモ
ナイ.

併シナガラ係數ガ自然数デナクテ, 實数ノ場合即チ *con-*
*tinuous parameter*ニ依存スル *Markoff process*ニ
於テ, 上ノコトハ如何ニナルカトイフコトハソレ程簡單デ
ハナイ.^{(*)2}

更ニ一般ニ可能ナ場合ガ実数ニヨリ標識付ケラレ, 且ツ
*continuous parameter*ニ依存スル *simple Markoff*

注意シテ)

$$P\left\{\sup_{0 \leq t \leq 1} |y_m(t)| \geq l\right\} \leq \frac{1}{l^2} \sum_{\Delta_m} E(b_{t_{i-1}}^2)(t_i - t_{i-1})$$

$m \rightarrow \infty$ トスレバ

$$P\left\{\sup_{0 \leq t \leq 1} \left| \int_0^t b_c d x_c \right| \geq l\right\} \leq \frac{1}{l^2} \int_0^1 E(b_c^2) dt$$

III. 微分方程式ト積分方程式

§9. 本章ト微分方程式

$$(1) dy_t = G(a(t, y_t), b(t, y_t))$$

ヲ

$$(2) y_0 = c$$

ト初期条件ノ下ニ解クコトヲ目的トスルニテ $G(a, b)$ ノ平均値 a , 標準偏差 b ト Gauss 分布ヲ表ハス。

先ガ定理ヲ掲ケル。

定理 8.1. $0 \leq t \leq 1, -\infty < y < \infty$ トルトキ $a(t, y)$

$b(t, y)$ ガ何レニ t, y 関シテ連続ナリ。且チ

$$(3) |a(t, y) - a(t, y')| \leq A |y - y'|$$

$$|b(t, y) - b(t, y')| \leq B |y - y'|$$

ヲ満スルコトナリ帯敷 A, B ガ存在スルトスル。然ラバ x_c ヲ

Brownian motion トスルトキ

$$(4) y_t = c + \int_0^t a(\tau, y_\tau) d\tau + \int_0^t b(\tau, y_\tau) dx_\tau$$

ハ一ツ而シテ唯一ツノ解ヲ有シ、ソノ解ハ (1) ヲ満足スル。

§10. 上ノ積分方程式ノ解ノ存在証明 (逐次近似法)

Lemma 10.1. $\{a_t\}$ ガ可動不連続点ナキ stochastic process ナリ $E(a_t^2) \leq M(t)$ ト連続函数 $M(t)$ ガ存在スルトスル。

而ラバ

$$(1) E\left\{\left(\int_t^s a_c d\tau\right)^2\right\} \leq (s-t) \int_t^s M(\tau) d\tau$$

証明

$$(2) \left(\int_t^s a_c d\tau\right)^2 \leq (s-t) \int_t^s a_c^2 d\tau$$

$$E\left\{\left(\int_t^s a_c d\tau\right)^2\right\} \leq (s-t) E\left\{\int_t^s a_c^2 d\tau\right\} \leq (s-t) \int_t^s M(\tau) d\tau$$

最後ノ不等式ハ Lemma 5.1 ヲ用ヒテ証明スルコトガ出来ル。

横ニ $t = \tau$ ナリ、 $y_t^{(k)}$ ($k=1, 2, \dots$) ヲ順々ニ次ノ

関係式ヲ定義スル。

$$(3) y_t^{(0)} = c$$

$$(4) y_t^{(k)} = c + \int_0^t a(\tau, y_\tau^{(k-1)}) d\tau + \int_0^t b(\tau, y_\tau^{(k-1)}) dx_\tau$$

($k=1, 2, \dots$)

先ガ $y_t^{(k)}$ ($k=1, 2, \dots$) ガ任意ノ t ($0 \leq t \leq 1$)

2 Chain rule for ODE

- ODE (Ordinary Differential Equation):

$$\frac{dX_t}{dt} = b(X_t)$$

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Namely,

$$df(X_t) = f'(X_t)b(X_t)dt$$

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Namely,

$$df(X_t) = f'(X_t)b(X_t)dt$$

$$f(X_T) - f(X_0) = \int_0^T f'(X_t)b(X_t)dt$$

$$= \lim_{|\Delta(T)| \rightarrow 0} \sum_i f'(X_{t_{k_i}})b(X_{t_{k_i}})(t_{i+1} - t_i)$$

$$\Delta(T) := \{0 = t_0 < t_1, \dots < t_N = T, N = 1, 2, \dots\}, t_i \leq t_{k_i} \leq t_{i+1}$$

3 Itô's differential rule

- SDE (Stochastic Differential Equation):

$$dX_t = b(X_t)dt + \sigma(X_t)dW_t$$

W_t : a standard Brownian motion

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W_t : a standard Brownian motion

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$$df(X_t) = \left\{ f'(X_t)b(X_t) + \frac{1}{2}f''(X_t)\sigma(X_t)^2 \right\}dt + f'(X_t)\sigma(X_t)dW_t$$

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Why " $\frac{1}{2}f''(X_t)\sigma(X_t)^2$ " comes in ?

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Why " $\frac{1}{2}f''(X_t)\sigma(X_t)^2$ " comes in ?

- simplest case ($b \equiv 0$, $\sigma(x) = 1$)

$$(1) \quad df(W_t) = \frac{1}{2}f''(W_t)dt + f'(W_t)dW_t$$

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- W_t is not of finite variation:

$$\sup_{\Delta(T)} \sum_{i=0}^N |W_{t_{i+1}} - W_{t_i}| = \infty, \quad \text{a.s.}$$

$$\Delta(T) := \{0 = t_0 < t_1 < t_2, \dots, < t_N = T, N = 1, 2, \dots\}$$

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- dW_t cannot be defined in an ordinary sense
- Then, how define dW_t ?
- Hint! look at the quadratic variation

$$\lim_{|\Delta(T)| \rightarrow 0} \sum |W_{t_{i+1}} - W_{t_i}|^2 = T$$

5 Brownian motion W_t (cont'd)

Note that for $0 = t_0 < t_1 \cdots < t_{N+1} = T$

$$\begin{aligned} W_T^2 - W_0^2 &= \sum_{i=0}^N (W_{t_{i+1}}^2 - W_{t_i}^2) \\ &= \sum_{i=0}^N (W_{t_{i+1}} - W_{t_i})^2 + \sum_{i=0}^N 2W_{t_i} (W_{t_{i+1}} - W_{t_i}) \end{aligned}$$

5 Brownian motion W_t (cont'd)

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As $|\Delta(T)| \rightarrow 0$ we have

$$(5.1) \quad W_T^2 - W_0^2 = T + \int_0^T 2W_t dW_t$$

if we define

$$\int_0^T W_t dW_t := \lim_{|\Delta|(T) \rightarrow 0} \sum W_{t_i} (W_{t_{i+1}} - W_{t_i})$$

$$(5.1) \quad W_T^2 - W_0^2 = T + \int_0^T 2W_t dW_t$$

(5.1) means

$$(5.2) \quad f(W_T) - f(W_0) = \int_0^T \frac{1}{2} f''(W_t) dt + \int_0^T f'(W_t) dW_t$$

for $f(x) = x^2$ since $f'(x) = 2x$, $f''(x) = 2$.

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For general f (5.2) can be seen by using Taylor expansion

$$\begin{aligned} f(W_{t_{i+1}}) - f(W_{t_i}) &= f'(W_{t_i})(W_{t_{i+1}} - W_{t_i}) \\ &\quad + \frac{1}{2} f''(W_{t_i})(W_{t_{i+1}} - W_{t_i})^2 + o(|W_{t_{i+1}} - W_{t_i}|^2) \end{aligned}$$

6 Brownian motion W_t (cont'd)

$$\begin{aligned}\sum_i W_{t_{i+1}} (W_{t_{i+1}} - W_{t_i}) &= \sum_i (W_{t_{i+1}} - W_{t_i})(W_{t_{i+1}} - W_{t_i}) \\ &\quad + \sum_i W_{t_i} (W_{t_{i+1}} - W_{t_i})\end{aligned}$$

implies

$$\lim_{|\Delta(T)| \rightarrow 0} \sum_i W_{t_{i+1}} (W_{t_{i+1}} - W_{t_i}) = T + \int_0^T W_t dW_t$$

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implies

$$\lim_{|\Delta(T)| \rightarrow 0} \sum_i W_{t_{i+1}} (W_{t_{i+1}} - W_{t_i}) = T + \int_0^T W_t dW_t$$

By defining

$$\int_0^T W_t \circ dW_t := \lim_{|\Delta(T)| \rightarrow 0} \sum_i \frac{W_{t_{i+1}} + W_{t_i}}{2} (W_{t_{i+1}} - W_{t_i})$$

we see that

$$\lim_{|\Delta(T)| \rightarrow 0} \sum_i W_{t_{i+1}} (W_{t_{i+1}} - W_{t_i}) = \int_0^T W_t \circ dW_t + \frac{1}{2}T$$

7 Definition of stochastic integral

Itô integral

For predictable $\varphi(t, \omega)$

$$\int_0^T \varphi(t, \omega) dW_t := \lim_{|\Delta(T)| \rightarrow 0} \sum_i \varphi(t_i, \omega) (W_{t_{i+1}} - W_{t_i})$$

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Stratonovich integral

For semimartingale $Y(t, \omega)$

$$\begin{aligned} \int_0^T Y(t, \omega) \circ dW_t &:= \lim_{|\Delta(T)| \rightarrow 0} \sum_i \frac{Y(t_{i+1}) + Y(t_i)}{2} (W_{t_{i+1}} - W_{t_i}) \\ &= \lim_{|\Delta(T)| \rightarrow 0} \sum_i \frac{2Y(t_i) + Y(t_{i+1}) - Y(t_i)}{2} (W_{t_{i+1}} - W_{t_i}) \\ &= \int_0^T Y_t dW_t + \frac{1}{2} \langle Y, W \rangle_t \end{aligned}$$

8 Itô's differential rule (II)

- Semi-martingale:

$$X_t = X_0 + \int_0^t b(X_s)ds + \int_0^t \sigma(X_s)dW_s \equiv X_0 + A_t + M_t$$

8 Itô's differential rule (II)

- Semi-martingale:

$$X_t = X_0 + \int_0^t b(X_s)ds + \int_0^t \sigma(X_s)dW_s \equiv X_0 + A_t + M_t$$

- Itô's differential rule for a semi-martingale:

$$\begin{aligned} f(X_t) - f(X_0) &= \int_0^t \left\{ f'(X_s)b(X_s) + \frac{1}{2}f''(X_s)\sigma(X_s)^2 \right\} ds \\ &\quad + \int_0^t f'(X_s)\sigma(X_s)dW_s \\ &= \int_0^t f'(X_s)dX_s + \frac{1}{2} \int_0^t f''(X_s)d\langle M, M \rangle_s \\ &= \int_0^t f'(X_s) \circ dX_s \end{aligned}$$

$$\langle M, M \rangle_t = \int_0^t \sigma(X_s)^2 ds, \quad \langle f'(X.), X. \rangle_t = \int_0^t f''(X_s)\sigma(X_s)^2 ds$$

9 Stochastic exponential

Let

$$Y_t = e^{\int_0^t \sigma(X_s) dW_s} \equiv e^{M_t}, \quad f(x) = e^x$$

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Then, by Itô's formula we have

$$df(M_t) = dY_t = Y_t \sigma(X_t) dW_t + \frac{Y_t}{2} \sigma(X_t)^2 dt,$$

namely,
$$\frac{dY_t}{Y_t} = \sigma(X_t) dW_t + \frac{1}{2} \sigma(X_t)^2 dt$$

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$$df(M_t) = dY_t = Y_t \sigma(X_t) dW_t + Y_t \frac{1}{2} \sigma(X_t)^2 dt,$$

namely,
$$\frac{dY_t}{Y_t} = \sigma(X_t) dW_t + \frac{1}{2} \sigma(X_t)^2 dt$$

Thus we see that

$$Z_t = e^{\int_0^t \sigma(X_t) dW_t - \frac{1}{2} \int_0^t \sigma(X_t)^2 dt}$$

satisfies

$$\frac{dZ_t}{Z_t} = \sigma(X_t) dW_t$$

- Differential equations determining a Markoff processes, Journal of Pan-Japan mathematical colloquium No.1077(1942) (original Japanese: Zenkoku Sizyo Sugaku Danwakai-si)
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10 Kolmogorov equation

- SDE

$$(10.1) \quad dX_t = b(X_t)dt + \sigma(X_t)dW_t, \quad X_0 = x$$

X_t^x : a solution to (10.1)

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$$(10.1) \quad dX_t = b(X_t)dt + \sigma(X_t)dW_t, \quad X_0 = x$$

X_t^x : a solution to (10.1)

- Itô's differential rule for $u = u(t, x)$

$$\begin{aligned} du(t, X_t^x) &= \left\{ \frac{\partial u}{\partial t}(t, X_t^x) + \frac{\partial u}{\partial x}(t, X_t^x)b(X_t^x) + \frac{1}{2} \frac{\partial^2 u}{\partial x^2}(t, X_t^x)\sigma(X_t^x)^2 \right\} dt \\ &\quad + \frac{\partial u}{\partial x}(t, X_t^x)\sigma(X_t^x)dW_t \\ &= \left\{ \frac{\partial u}{\partial t}(t, X_t^x) + Lu(t, X_t) \right\} dt + \frac{\partial u}{\partial x}(t, X_t^x)\sigma(X_t^x)dW_t \end{aligned}$$

where

$$Lu(t, x) := b(x) \frac{\partial u}{\partial x}(t, x) + \frac{1}{2} \sigma(x)^2 \frac{\partial^2 u}{\partial x^2}(t, x)$$

11 Kolmogorov equation (cont'd)

Set $u(s, x) := E[g(X_{T-s}^x)]$. If it is smooth, then it satisfies

$$\begin{cases} \frac{\partial u}{\partial t}(s, x) + Lu(s, x) = 0, & 0 \leq s < T \\ u(T, x) = g(x) \end{cases}$$

11 Kolmogorov equation (cont'd)

Set $u(s, x) := E[g(X_{T-s}^x)]$. If it is smooth, then it satisfies

$$\begin{cases} \frac{\partial u}{\partial t}(s, x) + Lu(s, x) = 0, & 0 \leq s < T \\ u(T, x) = g(x) \end{cases}$$

Because Itô's formula applied to $u(t, x)$ gives us

$$\begin{aligned} u(t, X_{t-s}^x) - u(s, x) &= \int_0^{t-s} \left\{ \left(\frac{\partial u}{\partial t} + Lu \right)(s+r, X_r^x) \right\} dr \\ &\quad + \int_0^{t-s} \frac{\partial u}{\partial x}(t, X_t^x) \sigma(X_t^x) dW_t \end{aligned}$$

and Markov property implies

$$\begin{aligned} u(s, x) &= E[g(X_{T-s}^x)] = E[E[g(X_{T-s}^x) | X_{t-s}^x]] \\ &= E[E[g(X_{T-t}^y) |_{y=X_{t-s}^x} | X_{t-s}^x]] = E[u(t, X_{t-s}^x)] \end{aligned}$$

12 Optimal portfolio selection

Market model

Price of a riskless asset:

$$dS_t^0 = S_t^0 r(X_t) dt, \quad S_0^0 = 1$$

Price of a risky asset:

$$dS_t^1 = S_t^1 \left\{ \alpha(X_t) dt + \sum_{j=1}^2 \sigma_j(X_t) dW_t^j \right\}, \quad S_0^1 = s_0$$

Economic factor:

$$dX_t = \beta(X_t) dt + \sum_{J=1}^2 \lambda_J(X_t) dW_t^J, \quad X_0 = x$$

Wealth process

$$V_t = N_t^0 S_t^0 + N_t^1 S_t^1$$

N_t^i , $i = 0, 1$: numbers of share invested to i -th security S_t^i

Wealth process

$$V_t := N_t^0 S_t^0 + N_t^1 S_t^1$$

N_t^i , $i = 0, 1$: numbers of share invested to i -th security S_t^i

Portfolio proportion

$$h_t^i := \frac{N_t^i S_t^i}{V_t}, \quad i = 0, 1, \quad h_t^0 + h_t^1 = 1$$

Denote h_t^1 by h_t in what follows since $h_t^0 = 1 - h_t^1$.

Wealth process

$$V_t := N_t^0 S_t^0 + N_t^1 S_t^1$$

N_t^i , $i = 0, 1$: numbers of share invested to i -th security S_t^i

Portfolio proportion

$$h_t^i := \frac{N_t^i S_t^i}{V_t}, \quad i = 0, 1, \quad h_t^0 + h_t^1 = 1$$

Denote h_t^1 by h_t in what follows since $h_t^0 = 1 - h_t^1$.

Dynamics of the wealth process

$$dV_t = V_t \{r(X_t) + h_t(\alpha(X_t) - r(X_t))\} dt + V_t h_t \sigma(X_t) dW_t,$$

$$V_0 = v_0$$

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$$V_0 = v_0$$

Dynamics of the wealth process

$$dV_t = V_t\{r(X_t) + h_t(\alpha(X_t) - r(X_t))\}dt + V_t h_t \sigma(X_t) dW_t,$$

$$V_0 = v_0$$

Since $\langle V \rangle_t = \int_0^t V_s^2 h_s^2 \sigma \sigma^*(X_s) ds$ we have by Itô's formula

$$\begin{aligned} dV_t^\gamma &= \gamma V_t^{\gamma-1} dV_t + \frac{1}{2} \gamma(\gamma-1) V_t^{\gamma-2} d\langle V \rangle_t \\ &= \gamma V_t^\gamma [\{r(X_t) + h_t(\alpha(X_t) - r(X_t))\}dt + h_t \sigma(X_t) dW_t] \\ &\quad + \frac{1}{2} \gamma(\gamma-1) V_t^\gamma h_t^2 \sigma \sigma^* dt \end{aligned}$$

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Therefore we have

$$\begin{aligned} \frac{dV_t^\gamma}{V_t^\gamma} &= \gamma \{r(X_t) + h_t(\alpha(X_t) - r(X_t)) + \frac{1}{2}(\gamma-1) h_t^2 \sigma \sigma^*(X_t)\} dt \\ &\quad + \gamma h_t \sigma(X_t) dW_t \end{aligned}$$

Once more again by Itô's formula we have

$$V_T^\gamma = v_0^\gamma e^\gamma \int_0^T \eta(X_s, h_s) ds + \gamma \int_0^T h_s \sigma(X_s) dW_s - \frac{\gamma^2}{2} \int_0^T h_s^2 \sigma \sigma^*(X_s) ds$$

where

$$\eta(x, h) = -\frac{1-\gamma}{2} h^2 \sigma \sigma^*(x) + h(\alpha(x) - r(x)) + r(x)$$

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Minimizing

$$\log E[V_T(h)^\gamma]$$

by selecting h amounts to minimizing

$$\log E\left[v_0^\gamma e^\gamma \int_0^T \eta(X_s, h_s) ds + \gamma \int_0^T h_s \sigma(X_s) dW_s - \frac{\gamma^2}{2} \int_0^T h_s^2 \sigma \sigma^*(X_s) ds\right]$$

Utility maximization and risk-sensitive portfolio optimization

Power utility maximization:

$$\sup_h \frac{1}{\gamma} E[V_T(h)^\gamma], \quad \gamma < 0$$

is equivalent to risk-sensitive portfolio optimization:

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and so

$$\inf_h \log E\left[v_0^\gamma e^{\gamma \int_0^T \eta(X_s, h_s) ds + \gamma \int_0^T h_s \sigma(X_s) dW_s - \frac{\gamma^2}{2} \int_0^T h_s^2 \sigma \sigma^*(X_s) ds}\right]$$

Through change of measure

$$\frac{dP^h}{dP} \Big|_{\mathcal{F}_T} = e^{\gamma \int_0^T h_s \sigma(X_s) dW_s - \frac{\gamma^2}{2} \int_0^T h_s^2 \sigma \sigma^*(X_s) ds},$$

$$\inf_h \log E[v_0^\gamma e^{\gamma \int_0^T \eta(X_s, h_s) ds + \gamma \int_0^T h_s \sigma(X_s) dW_s - \frac{\gamma^2}{2} \int_0^T h_s^2 \sigma \sigma^*(X_s) ds}]$$

turns out to be

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whose H-J-B (Hamilton-Jacobi-Bellman) equation is
(12.1)

$$\begin{aligned} \frac{\partial v}{\partial t} + \frac{1}{2} \lambda \lambda^* \frac{\partial^2 v}{\partial x^2} + \frac{1}{2} \lambda \lambda^* \left| \frac{\partial v}{\partial x} \right|^2 \\ + \inf_h \{ [\beta(x) + \gamma \lambda \sigma^*(x) h] \frac{\partial v}{\partial x} + \gamma \eta(x, h) \} = 0 \end{aligned}$$

$$v(T, x) = \gamma \log v_0$$

Analysis of the H-J-B equation (12.1) gives us optimal portfolio strategy

$$\hat{h}(t, X_t)$$

under suitable conditions, where

$$\hat{h}(t, x) = \frac{1}{1 - \gamma} (\sigma \sigma^*(x))^{-1} (\alpha(x) - r(x)) + \sigma \lambda^*(x) \frac{\partial v}{\partial x}(t, x).$$

and $\frac{\partial v}{\partial x}(t, x)$ is the solution of the H-J-B equation.

Namely, $\hat{h}(t, X_t)$ attains the minimum of

$$\log E^h \left[v_0^\gamma e^{\gamma \int_0^T \eta(X_s, h_s) ds} \right], \quad \gamma < 0$$

which is equivalent to utility maximization.

Down-side risk minimization

Consider a down-side risk minimization on a long term:

$$J(\kappa) := \inf_h \liminf_{T \rightarrow \infty} \frac{1}{T} \log P\left(\frac{1}{T} \log V_T(h) \leq \kappa\right)$$

for a given target growth rate κ

and risk-sensitive portfolio optimization on a long term:

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We can show that

$$J(\kappa) = - \inf_{k \in (-\infty, \kappa]} \sup_{\gamma < 0} \{\gamma k - \chi(\gamma)\} = - \sup_{\gamma < 0} \{\gamma \kappa - \chi(\gamma)\}$$

for $\lim_{\gamma \rightarrow \infty} \chi'(\gamma) < \kappa < \chi'(0-)$ under suitable conditions.

Moreover, for given target growth rate κ we take $\gamma(\kappa)$ which attains

$$\sup_{\gamma < 0} \{\gamma\kappa - \chi(\gamma)\} = \gamma(\kappa)\kappa - \chi(\gamma(\kappa))$$

and choose an optimal strategy $\hat{h}_t^{\gamma(\kappa)}$ which minimize

$$\liminf_{T \rightarrow \infty} \frac{1}{T} \log E[V_T(h)^\gamma].$$

Then, this $\hat{h}_t^{\gamma(\kappa)}$ attains the minimum of

$$\liminf_{T \rightarrow \infty} \frac{1}{T} \log P\left(\frac{1}{T} \log V_T(h) \leq \kappa\right)$$

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we can obtain $\hat{h}_t^{\gamma(\kappa)}$ through analysis of H-J-B equation of ergodic type corresponding to risk-sensitive portfolio optimization:

$$\chi(\gamma) := \inf_h \liminf_{T \rightarrow \infty} \frac{1}{T} \log E[V_T(h)^\gamma].$$

Thank you for your attention !